

1. Consider a coin with $P(H) = p$ and $P(T) = q$. Flip the coin until the first head shows up and let X be the number of flips you made. The probability mass function and support of this *geometric random variable* are given by

$$P(X = k) = q^{k-1}p \quad \text{and} \quad S_X = \{1, 2, 3, \dots\}.$$

(a) Use the geometric series $1 + q + q^2 + \dots = (1 - q)^{-1}$ to show that

$$\sum_{k \in S_X} P(X = k) = 1.$$

(b) Differentiate the geometric series to get $0 + 1 + 2q + 3q^2 + \dots = (1 - q)^{-2}$ and use this series to show that

$$E[X] = \sum_{k \in S_X} k \cdot P(X = k) = \frac{1}{p}.$$

(c) Application: Start rolling a fair 6-sided die. On average, how long do you have to wait until you see “1” for the first time?

Throughout I will assume that $0 < q = 1 - p < 1$. (The cases $p \in \{0, 1\}$ are slightly different.) For (a) we have

$$\begin{aligned} \sum_{k \in S_X} P(X = k) &= \sum_{k=1}^{\infty} P(X = k) \\ &= \sum_{k=1}^{\infty} q^{k-1}p \\ &= p + qp + q^2p + q^3p + \dots \\ &= p(1 + q + q^2 + q^3 + \dots) \\ &= p/(1 - q) \\ &= p/p \\ &= 1. \end{aligned}$$

In words: The total probability is 1. Then for (b) we have

$$\begin{aligned} E[X] &= \sum_{k \in S_X} k \cdot P(X = k) = \sum_{k=1}^{\infty} k \cdot P(X = k) \\ &= \sum_{k=1}^{\infty} kq^{k-1}p \\ &= p + 2qp + 3q^2p + 4q^3p + \dots \\ &= p(1 + 2q + 3q^2 + 4q^3 + \dots) \\ &= p/(1 - q)^2 \\ &= p/p^2 \\ &= 1/p. \end{aligned}$$

In words: We expect to see the first H on the $(1/p)$ -th flip of the coin.

(c) Application: For example, we can think of a fair six-sided die as a strange coin where $H = \{\text{we get 1}\}$ and $T = \{\text{we don't get 1}\}$, so that $P(H) = p = 1/6$. Let X be the number of rolls until we see "1." Then by part (b) we have

$$E[\text{number of rolls until our first 1}] = E[X] = \frac{1}{p} = \frac{1}{1/6} = 6.$$

That makes sense.

2. There are 2 red balls and 4 green balls in an urn. Suppose you grab 3 balls without replacement and let X be the number of red balls you get.

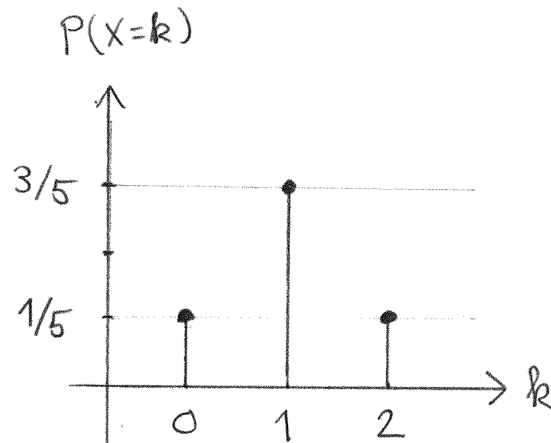
- What is the support of this random variable?
- Draw a picture of the probability mass function $f_X(k) = P(X = k)$.
- Compute the expected value $E[X]$. Does the answer make sense?

(a) The support is $S_X = \{0, 1, 2\}$.

(b) This X is a hypergeometric random variable with pmf given by the following table:

k	0	1	2
$P(X = k)$	$\frac{\binom{2}{0}\binom{4}{3}}{\binom{6}{3}} = \frac{4}{20}$	$\frac{\binom{2}{1}\binom{4}{2}}{\binom{6}{3}} = \frac{12}{20}$	$\frac{\binom{2}{2}\binom{4}{1}}{\binom{6}{3}} = \frac{4}{20}$

Here is the line graph:



(c) We have

$$E[X] = 0 \cdot P(X = 0) + 1 \cdot P(X = 1) + 2 \cdot P(X = 2) = 0 \cdot \frac{1}{5} + 1 \cdot \frac{3}{5} + 2 \cdot \frac{1}{5} = \frac{5}{5} = 1.$$

This answer makes sense for two reasons:

- The line graph is symmetric about $k = 1$, so $k = 1$ is the center of mass.
- One third of the balls in the urn are red. If we grab three balls then we expect one third of them (i.e., one ball) to be red.

3. Roll a pair of fair 6-sided dice and consider the following random variables:

X = the number that shows up on the first roll,

Y = the number that shows up on the second roll.

- (a) Write down all elements of the sample space S .
- (b) Compute the probability mass function for the sum $f_{X+Y}(k) = P(X + Y = k)$ and draw the probability histogram.
- (c) Compute the expected value $E[X + Y]$ in two different ways.
- (d) Compute the probability mass function for the difference $f_{X-Y}(k) = P(X - Y = k)$ and draw the probability histogram.
- (e) Compute the expected value $E[X - Y]$ in two different ways.
- (f) Compute the probability mass function for the absolute value of the difference

$$f_{|X-Y|}(k) = P(|X - Y| = k)$$

and draw the probability histogram.

- (g) Compute the expected value $E[|X - Y|]$. This time there is only one way to do it.

- (a) I will assume that the dice are ordered.¹ Here is the sample space:

$$S = \left\{ \begin{array}{cccccc} 11 & 12 & 13 & 14 & 15 & 16 \\ 21 & 22 & 23 & 24 & 25 & 26 \\ 31 & 32 & 33 & 34 & 35 & 36 \\ 41 & 42 & 43 & 44 & 45 & 46 \\ 51 & 52 & 53 & 54 & 55 & 56 \\ 61 & 62 & 63 & 64 & 65 & 66 \end{array} \right\}$$

- (b) To compute the probabilities $P(X + Y = k)$ we circle the events $\{X + Y = k\}$:

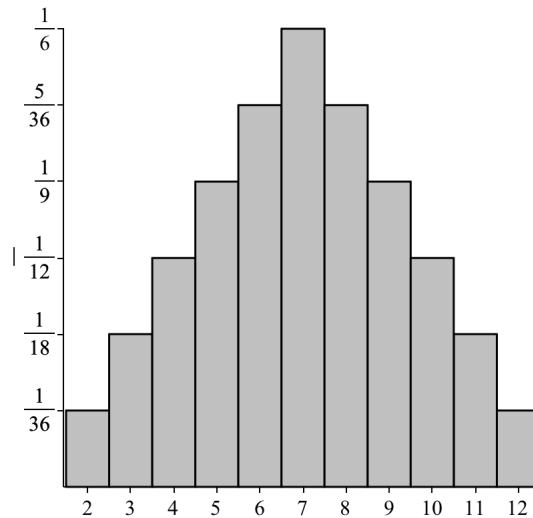
$$X + Y = \left\{ \begin{array}{cccccc} 2 & 3 & 4 & 5 & 6 & 7 \\ \textcircled{11} & \textcircled{12} & \textcircled{13} & \textcircled{14} & \textcircled{15} & \textcircled{16} \\ 21 & 22 & 23 & 24 & 25 & 26 \\ 31 & 32 & 33 & 34 & 35 & 36 \\ 41 & 42 & 43 & 44 & 45 & 46 \\ 51 & 52 & 53 & 54 & 55 & 56 \\ \textcircled{61} & \textcircled{62} & \textcircled{63} & \textcircled{64} & \textcircled{65} & \textcircled{66} \end{array} \right\}$$

¹Unordered dice makes the problem much harder.

And then we count them up:

k	2	3	4	5	6	7	8	9	10	11	12
$P(X + Y = k)$	$\frac{1}{36}$	$\frac{2}{36}$	$\frac{3}{36}$	$\frac{4}{36}$	$\frac{5}{36}$	$\frac{6}{36}$	$\frac{5}{36}$	$\frac{4}{36}$	$\frac{3}{36}$	$\frac{2}{36}$	$\frac{1}{36}$

Here is the probability histogram:



(c) We can compute the expected value directly from the formula:

$$\begin{aligned}
 E[X + Y] &= \sum_{k=2}^{12} k \cdot P(X + Y = k) \\
 &= 2 \cdot \frac{1}{36} + 3 \cdot \frac{2}{36} + \cdots + 11 \cdot \frac{2}{36} + 12 \cdot \frac{1}{36} = \frac{252}{36} = 7.
 \end{aligned}$$

Or we can use linearity. Recall from class that the average outcome in one roll of a fair die is $E[X] = E[Y] = 3.5$. Thus we have

$$E[X + Y] = E[X] + E[Y] = 3.5 + 3.5 = 7.$$

(d) To compute the probabilities $P(X - Y = k)$ we circle the events $\{X - Y = k\}$:

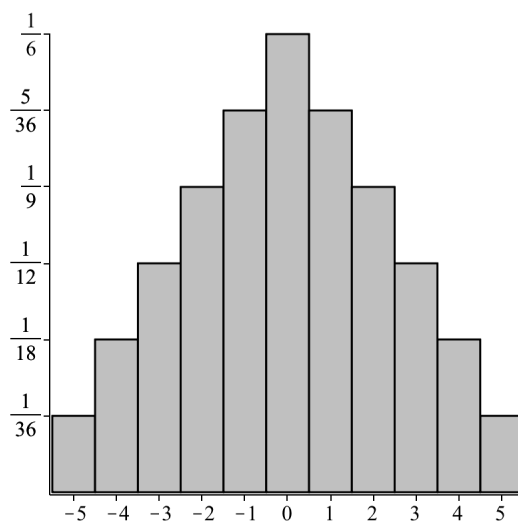
$X - Y =$

}	11	12	13	14	15	16	-5
	21	22	23	24	25	26	-4
	31	32	33	34	35	36	-3
	41	42	43	44	45	46	-2
	51	52	53	54	55	56	-1
	61	62	63	64	65	66	}
		5	4	3	2	1	0

And then we count them up:

k	-5	-4	-3	-2	-1	0	1	2	3	4	5
$P(X - Y = k)$	$\frac{1}{36}$	$\frac{2}{36}$	$\frac{3}{36}$	$\frac{4}{36}$	$\frac{5}{36}$	$\frac{6}{36}$	$\frac{5}{36}$	$\frac{4}{36}$	$\frac{3}{36}$	$\frac{2}{36}$	$\frac{1}{36}$

Here is the probability histogram:



Note that the histogram looks the same, just shifted 7 units to the left. That surprised me.

(e) We can compute the expected value directly from the formula:

$$\begin{aligned}
 E[X - Y] &= \sum_{k=-5}^5 k \cdot P(X - Y = k) \\
 &= -5 \cdot \frac{1}{36} - 4 \cdot \frac{2}{36} - \dots + 4 \cdot \frac{2}{36} + 5 \cdot \frac{1}{36} = \frac{0}{36} = 0.
 \end{aligned}$$

Or we can use linearity:

$$E[X - Y] = E[X] - E[Y] = 3.5 - 3.5 = 0.$$

(f) When we compute $P(|X - Y| = k)$, some of the values from (d) turn positive:

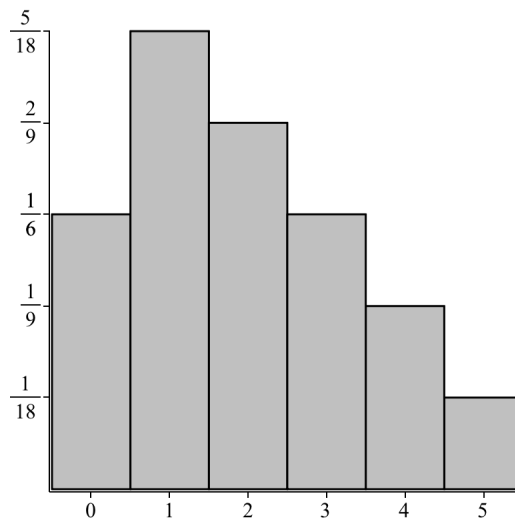
$$|X - Y| =$$

}	11	12	13	14	15	16	+5
	21	22	23	24	25	26	+4
	31	32	33	34	35	36	+3
	41	42	43	44	45	46	+2
	51	52	53	54	55	56	+1
	61	62	63	64	65	66	}
		5	4	3	2	1	0

Then we add them up:

k	0	1	2	3	4	5
$P(X - Y = k)$	$\frac{6}{36}$	$\frac{10}{36}$	$\frac{8}{36}$	$\frac{6}{36}$	$\frac{4}{36}$	$\frac{2}{36}$

And draw the histogram:



(g) We can compute the expected value directly from the formula:

$$\begin{aligned}
 E[|X - Y|] &= \sum_{k=0}^5 k \cdot P(|X - Y| = k) \\
 &= 0 \cdot \frac{6}{36} + 1 \cdot \frac{10}{36} + \dots + 4 \cdot \frac{4}{36} + 5 \cdot \frac{2}{36} = \frac{70}{36} = 1.94.
 \end{aligned}$$

This time the linearity of expectation doesn't help because the absolute value is not nice.

4. Let X be a random variable satisfying

$$E[X + 1] = 3 \quad \text{and} \quad E[(X + 1)^2] = 10.$$

Use this information to compute the following:

$$\text{Var}(X + 1), \quad E[X], \quad E[X^2] \quad \text{and} \quad \text{Var}(X).$$

Solution: By definition we have

$$\text{Var}(X + 1) = E[(X + 1)^2] - E[X + 1]^2 = 10 - 3^2 = 1.$$

Then using linearity of expectation gives

$$E[X + 1] = 3$$

$$E[X] + E[1] = 3$$

$$E[X] + 1 = 3$$

$$E[X] = 2$$

and

$$E[(X + 1)^2] = 10$$

$$E[X^2 + 2X + 1] = 10$$

$$E[X^2] + 2E[X] + 1 = 10$$

$$E[X^2] + 2(2) + 1 = 10$$

$$E[X^2] = 5.$$

Finally, we have

$$\text{Var}(X) = E[X^2] - E[X]^2 = 5 - 2^2 = 1.$$

Note that this agrees with the general fact $\text{Var}(X + \alpha) = \text{Var}(X)$ when α is constant.

5. Let X be a random variable with mean $E[X] = \mu$ and variance $\text{Var}(X) = \sigma^2 \neq 0$. Compute the mean and variance of the random variable Y defined by

$$Y = \frac{X - \mu}{\sigma}.$$

Solution: For the expected value we have

$$E[Y] = E\left[\frac{X - \mu}{\sigma}\right] = \frac{1}{\sigma}E[X - \mu] = \frac{1}{\sigma}(E[X] - \mu) = \frac{1}{\sigma}(\mu - \mu) = 0.$$

For the variance we have

$$\begin{aligned} E[Y] &= E[Y^2] - \cancel{E[Y]^2} \\ &= E[Y^2] \\ &= E\left[\frac{(X - \mu)^2}{\sigma^2}\right] \\ &= \frac{1}{\sigma^2}E[(X - \mu)^2] \\ &= \frac{1}{\sigma^2}\text{Var}(X) = \frac{1}{\sigma^2} \cdot \sigma^2 = 1. \end{aligned}$$

6. Let X be the number of strangers you must talk to until you find someone who shares your birthday. (Assume that each day of the year is equally likely and ignore February 29.)

- (a) Find the probability mass function $P(X = k)$.
 (b) Find the expected value $\mu = E[X]$.
 (c) Find the *cumulative mass function* $P(X \leq k)$. Hint: If X is a geometric random variable with pmf $P(X = k) = q^{k-1}p$, use the geometric series to show that

$$P(X \leq k) = 1 - P(X > k) = 1 - \sum_{i=k+1}^{\infty} q^{i-1}p = 1 - q^k.$$

- (d) Use part (c) to find the probability $P(\mu - 50 \leq X \leq \mu + 50)$ that X falls within ± 50 of the expected value. Hint:

$$P(\mu - 50 \leq X \leq \mu + 50) = P(X \leq \mu + 50) - P(X \leq \mu - 50 - 1).$$

(a) We can think of each stranger as a coin flip where “heads” means “they have the same birthday as you.” Then X is a geometric random variable with $P(H) = p = 1/365$ and $P(T) = q = 1 - p = 364/365$. From Problem 1 we know that

$$P(X) = q^{k-1}p = \left(\frac{364}{365}\right)^{k-1} \left(\frac{1}{365}\right) = \frac{364^{k-1}}{365^k}.$$

(b) From Problem 1 we also know that

$$E[X] = \frac{1}{p} = \frac{1}{1/365} = 365.$$

That is, on average you will need to speak to 365 strangers until you find someone who shares your birthday. That makes sense.

(c) Note that $P(X \leq k) = 1 - P(X > k)$ and

$$\begin{aligned} P(X > k) &= P(X = k + 1) + P(X = k + 2) + P(X = k + 3) + \dots \\ &= q^k p + q^{k+1} p + q^{k+2} p + \dots \\ &= q^k p (1 + q + q^2 + \dots) \\ &= q^k p \cdot \frac{1}{1 - q} = q^k p \cdot \frac{1}{p} = q^k. \end{aligned}$$

So we conclude that $P(X \leq k) = 1 - q^k$.

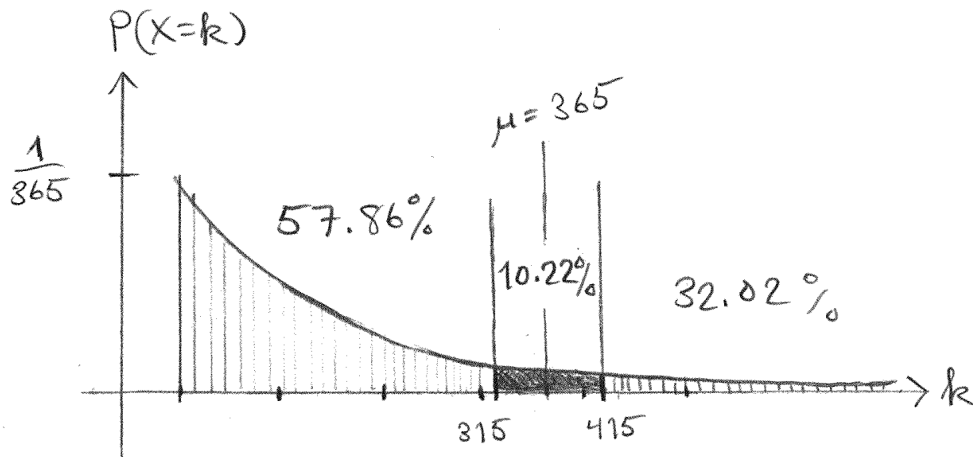
(d) Continuing from part (c), we have for any whole numbers k and ℓ that

$$P(k \leq X \leq \ell) = P(X \leq \ell) - P(X \leq k - 1) = (1 - q^\ell) - (1 - q^{k-1}) = q^{k-1} - q^\ell.$$

In particular, we see that

$$P(315 \leq X \leq 415) = q^{314} - q^{415} = \left(\frac{364}{365}\right)^{314} - \left(\frac{364}{365}\right)^{415} = 10.22\%.$$

In other words, there is a 10.22% chance that you will need to ask between 315 and 415 people until you find someone who shares your birthday. Here is a picture of the pmf (not to scale, obviously):



7. I am running a lottery. I will sell 10 tickets, each for a price of \$1. The person who buys the winning ticket will receive a cash prize of \$5.

- If you buy one ticket, what is the expected value of your profit?
- If you buy two tickets, what is the expected value of your profit?
- If you buy n tickets ($0 \leq n \leq 10$), what is the expected value of your profit? Which value of n maximizes your expected profit?

[Remark: Profit equals prize money minus cost of the tickets.]

(a) Suppose you buy one ticket and let X be your profit. If you buy a losing ticket then your profit is $X = -1$ dollar and if you buy the winning ticket then your profit is $X = -1 + 5 = 4$ dollars. The probability of getting the winning ticket is $1/10$, so here is the pmf of X :

k	-1	$+4$
$P(X = k)$	$\frac{9}{10}$	$\frac{1}{10}$

Your expected profit is $E[X] = (-1)(9/10) + 4(1/10) = -5/10 = -0.5$ dollars.

(b) Let X be your profit from the purchase of two tickets. If both tickets are losers then $X = -2$ and if one ticket is a winner then $X = -2 + 5 = 3$. The number of ways to choose 2 out of 10 tickets is $\binom{10}{2} = 45$ and the number of way to choose 2 losing tickets out of 9 is $\binom{9}{2} = 36$. The number of ways to choose 1 winning ticket and 1 losing ticket is $\binom{1}{1}\binom{9}{1} = 9$. So here is the pmf of X :

k	-2	$+3$
$P(X = k)$	$\frac{\binom{1}{0}\binom{9}{2}}{\binom{10}{2}} = \frac{36}{45} = \frac{8}{10}$	$\frac{\binom{1}{1}\binom{9}{1}}{\binom{10}{2}} = \frac{9}{45} = \frac{2}{10}$

Your expected profit is $E[X] = (-2)(8/10) + 3(2/10) = -10/10 = -1$ dollar.

(c) Let X be your profit from the purchase of n tickets (where $0 \leq n \leq 10$). Using a similar argument (and some algebraic manipulation) gives the pmf:

$$P(X = k) \left| \begin{array}{c|c|c} k & -n & 5-n \\ \hline \frac{\binom{1}{0}\binom{9}{n}}{\binom{10}{n}} = \frac{10-n}{10} & & \frac{\binom{1}{1}\binom{9}{1}}{\binom{10}{2}} = \frac{n}{10} \end{array} \right.$$

Your expected profit is $E[X] = (-n)(10-n)/10 + (5-n)(n/10) = -n/2$ dollars. You will maximize your profit by purchasing $n = 0$ lottery tickets. But there is an easier way.

Easier Solution: Suppose you purchase n tickets and let X_i be your profit from the i th ticket. From part (a) we know that $E[X_i] = -0.5$. Now let $X = X_1 + X_2 + \dots + X_n$ be your total profit. Using linearity of expectation gives

$$\begin{aligned} X &= X_1 + X_2 + \dots + X_n \\ E[X] &= E[X_1 + X_2 + \dots + X_n] \\ E[X] &= E[X_1] + E[X_2] + \dots + E[X_n] \\ &= -0.5 - 0.5 - \dots - 0.5 = n(-0.5) = -n/2. \end{aligned}$$

Note that linearity of expectation holds even though the random variables X_i are **not** independent of each other. That's pretty useful.

8. Consider a coin with $P(H) = p$ and $P(T) = q$. Flip the coin n times and let X be the number of heads you get. In this problem you will give a bad proof that $E[X] = np$.

- (a) Use the formula $\binom{n}{k} = \frac{n!}{k!(n-k)!}$ to show that $k\binom{n}{k} = n\binom{n-1}{k-1}$.
 (b) Complete the following computation:

$$\begin{aligned} E[X] &= \sum_{k=0}^n k \cdot P(X = k) \\ &= \sum_{k=1}^n k \cdot P(X = k) \\ &= \sum_{k=1}^n k \binom{n}{k} p^k q^{n-k} \\ &= \sum_{k=1}^n k \binom{n}{k} p^k q^{n-k} \\ &= \sum_{k=1}^n n \binom{n-1}{k-1} p^k q^{n-k} \\ &= \dots \end{aligned}$$

(a) Note that

$$n \binom{n-1}{k-1} = n \cdot \frac{(n-1)!}{(k-1)![(n-1)-(k-1)]!} = \frac{n(n-1)!}{(k-1)(n-k)!} = \frac{n!}{(k-1)!(n-k)!}$$

and

$$k \binom{n}{k} = k \cdot \frac{n!}{k!(n-k)!} = k \cdot \frac{n!}{k(k-1)!(n-k)!} = \frac{n!}{(k-1)!(n-k)!}.$$

Alternate Proof: Suppose you want to choose a k -member club from a classroom of n students, where one member of the club will serve as president. On the one hand, you can choose the club members in $\binom{n}{k}$ ways. Then there are k ways to choose the president. On the other hand, you can first choose a student to serve as club president. There are n ways to do this. Then there are $\binom{n-1}{k-1}$ ways to choose the remaining $k-1$ club members from the remaining $n-1$ students.

(b) Since $p + q = 1$, let me first note that

$$1 = 1^{n-1} = (p + q)^{n-1} = \sum_{\ell=0}^{n-1} \binom{n-1}{\ell} p^{\ell} q^{(n-1)-\ell}.$$

Then making the substitution $k = \ell + 1$ gives

$$\begin{aligned} E[X] &= \dots \\ &= \sum_{k=1}^n n \binom{n-1}{k-1} p^k q^{n-k} \\ &= n \sum_{k=1}^n \binom{n-1}{k-1} p^k q^{n-k} \\ &= np \left(\sum_{k=1}^n \binom{n-1}{k-1} p^{k-1} q^{n-k} \right) \\ &= np \left(\sum_{\ell=0}^{n-1} \binom{n-1}{(\ell+1)-1} p^{(\ell+1)-1} q^{n-(\ell+1)} \right) \\ &= np \left(\sum_{\ell=0}^{n-1} \binom{n-1}{\ell} p^{\ell} q^{(n-1)-\ell} \right) \\ &= np(1) \\ &= np, \end{aligned}$$

as desired.

Remark: That was certainly a bad way to solve the problem.